Process evaluation and causal mediation analysis using Mplus

Dr Richard Emsley
Biostatistics, Health Sciences – Methodology, School of Community Based Medicine, The University of Manchester, Manchester, UK

www.medicine.manchester.ac.uk/staff/RichardEmsley
richard.emsley@manchester.ac.uk

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  ➢ Richard Emsley (MRC Fellow), Graham Dunn.

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  ➢ Graham Dunn (PI), Richard Emsley, Linda Davies, Jonathan Green, Andrew Pickles, Chris Roberts, Ian White & Frank Windmeijer with Hanhua Liu.
Contents

1. Discussion of mediation and statistical mediation analysis

2. Causal mediation parameters

3. Process evaluation (using Mplus)

4. Causal mediation analysis (using Mplus)

5. Conclusions
Mediation analysis and causal inference...

“Mediation analysis is a form of causal analysis...all too often persons conducting mediational analysis either do not realize that they are conducting causal analyses or they fail to justify the assumptions that they have made in their casual model.”

The basic underlying problem: estimating valid causal effects

U – the unmeasured confounders

Total effect = direct effect (γ) + indirect effect (α*β)
Solutions to unmeasured confounding

- We’ve proposed three solutions to analyse mediation allowing for unmeasured confounding:

  1. Measure and adjust for potential confounders (sounds obvious, not always done);

  2. Instrumental variables;

  3. Principal stratification.

Explained in detail in:
True and Incidental mediators

- **True mediators**: intermediate variables which test the mechanism/theory through which an intervention acts.

- Examples in psychological treatment trials:
  - **Do people jump to conclusions?** *(PRP trial)*  
    Does psychotherapy reduce jumping to conclusions which improves positive symptoms in psychosis?

  - **What is the concomitant substance abuse?** *(MIDAS trial)*  
    Does psychotherapy reduce cannabis use, which in turn leads to improvements in psychotic symptoms?

- **Incidental mediators**: variables measured post-randomisation that we may wish to rule out having a mediated effect.

  - **Use of concomitant medication** *(PROSPECT trial)*  
    Does psychotherapy improve compliance with medication which, in turn, leads to better outcome?
True and Incidental mediators

- What makes these variables ‘mediators’?
  - We are interested in all three pathways in the diagram, and the effect decomposition:

![Diagram](image)

- New requirements for mediation?
  1. Aim is to estimate the size of the indirect effect, and
  2. The mediator is measured in both arms.
Statistical mediation analysis

- Large literature on statistical mediation analysis, summarised by the recent monograph by David MacKinnon (2008).

- Further work by Kris Preacher and Andrew Hayes, developing SPSS macros for multiple mediators, moderated mediation/mediated moderation, longitudinal mediation models.

- Extensive use of structural equation modelling including Mplus examples.

- All (usually) based on the same implicit assumptions.
Characteristics of therapy: mediators or post-randomisation effect modifiers?

- Aspects involved in process of therapy that might explain differential treatment effects/effect heterogeneity.

  - **Compliance with allocated treatment**
    Does the participant turn up for any therapy? How many sessions does she attend?

  - **Quality of the therapeutic relationship**
    What is the strength of the therapeutic alliance?

  - **Fidelity of therapy**
    How close is the therapy to that described in the treatment manual? Is it a cognitive-behavioural intervention, for example, or merely emotional support?
Characteristics of therapy: mediators or post-randomisation effect modifiers?

- Why do I argue these aren’t true or incidental mediators?
  - Generally interested in some other causal question, such as how do they account for heterogeneity? Are they effect modifiers?

![Diagram showing relationships between Randomisation to CBT, Sessions, and Depression Score]
Characteristics of therapy: latent variables?

- Fidelity of therapy
- Components of therapy
- Quality of therapeutic relationship
- Therapeutic dose

It is plausible that these may only be measured in the therapy arm of a randomised trial.

For example, if the control arm has some form of treatment as usual which doesn’t contain an active ‘therapy’ on which they can be measured.
Psychosis Research Partnership – engagement in therapy example

Randomisation → Patient engagement in therapy → PANSS
Psychosis Research Partnership – engagement in therapy example

- **Full therapy**
  - Randomisation
  - ITT effect
  - PANSS

- **Partial therapy**
  - Randomisation
  - ITT effect
  - PANSS

- **No therapy**
  - Randomisation
  - ITT effect
  - PANSS
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Notation

- $Z_i$ – randomised group ($Z_i=1$ for treatment, 0 for controls).

- $X_i$ – baseline covariates.

- $Y_i$ – observed outcome.

- $M_i$ – intermediate outcome that is a putative mediator of the effects of treatment on outcome (either a quantitative measure or binary).

- $Y_i(0)$ – counterfactual treatment free outcome
Notation and counterfactuals

We define the following counterfactual outcomes:

\( M_i(z) \) – mediator with treatment \( Z=z \).

\( Y_i(z,m) \) – outcome with treatment \( Z=z \) and level of mediator \( M=m \).

\( Y_i(0) = Y_i(0, M_i(0)) \) – outcome if \( Z=0 \) with mediator \( M_i(0) \).

\( Y_i(1) = Y_i(1, M_i(1)) \) – outcome if \( Z=1 \) with mediator \( M_i(1) \).

In the control arm, \( Y_i = Y_i(0) \) and \( M_i = M_i(0) \), so \( M_i(0) \) and \( Y_i(0) \) are observed and \( M_i(1) \) and \( Y_i(1) \) are unobserved.

Similarly, in the treatment arm, \( M_i(0) \) and \( Y_i(0) \) are unobserved and \( M_i = M_i(1) \) and \( Y_i = Y_i(1) \) are observed.
Causal mediation definitions: direct and indirect effects

- **(Pure) natural direct effect**: $Y_i(1,M_i(0)) - Y_i(0,M_i(0))$
  - The direct effect of random allocation given $M(0)$, the `natural` level of the mediator

- **(Total) natural indirect effect**: $Y_i(1,M_i(1)) - Y_i(1,M_i(0))$
  - The effect of the change in mediator if randomised to receive treatment (i.e. $Z=1$).

- **Controlled direct effect**: $Y_i(1,m) - Y_i(0,m)$
  - Direct effect of randomisation on outcome at mediator level $m$.

- Total Effect = Natural direct effect + Natural indirect effect

(Pearl 2001; Robins & Greenland 1992).
An alternative approach for post-randomisation effect-modifiers

- When the intermediate variable (e.g. therapeutic alliance or treatment fidelity) is not observed in the control arm, we can also estimate a **principal stratum direct effect**: 

\[
\text{PSDE} = \mathbb{E}[Y(1)-Y(0) | M(1)=m]
\]

- This uses an approach called principal stratification.

- Key issue is to predict \(M(1)\) when \(Z=0\).
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What is principal stratification?

- It involves classifying subjects into classes which are defined by their joint potential responses of the intermediate variable to all possible random allocations.

- Rather than using the observed value of this intermediate variable (which may not be possible if it is not measured in the control group), it is more useful to consider the potential value if an individual were allocated to active treatment, which is observed in the treatment arm but unobserved in the control arm.

- These classes are known as principal strata which have the property that they are independent of treatment allocation and can be handled in the analysis in an analogous way to pre-randomisation variables.

Principal strata – therapeutic alliance example
Principal strata – therapeutic alliance example
Principal strata – model identification

- If the intermediate variable is only measured in the treatment condition (e.g. therapeutic alliance when no intervention offered in the control group), then we know the stratum membership for the treatment group.

- We need baseline data that will strongly predict class membership, and use this to predict class membership for the control group.

- Essentially this is just a finite mixture model.

- The key identifying assumption is that there are no treatment by covariate interactions which have an effect on the outcome, but which do have an effect on the intermediate variable.
Example: SoCRATES summary

- **SoCRATES** (*Study of Cognitive Re-Alignment therapy in Early Schizophrenia*) trial was designed to evaluate the effects of **cognitive behaviour therapy** and **supportive counselling** on the outcomes of patients after an early episode of schizophrenia.

- For our illustrative purposes, we ignore the distinction between CBT and SC, using a binary variable for treatment (CBT or SC, N=207) and control (TAU, N=102).

- Recruitment and randomisation was within **3 treatment centres**: Liverpool, Manchester and Nottinghamshire. Other baseline covariates include **logarithm of untreated psychosis** and **years of education**.

- Outcome was the **Positive and Negative Syndromes Schedule (PANSS)**, an interview-based scale for rating psychotic and non-psychotic symptoms ranging from 30 to 210 (high scores imply worse symptoms).

Lewis et al, BJP (2002); Tarrier et al BJP (2004); Dunn & Bentall, Stats in Medicine (2007); Emsley, Dunn and White, Stats Methods in Medial Research (2010).
Example: therapeutic alliance in SoCRATES

- Therapeutic alliance was measured at the 4th session of therapy, early in the time-course of the intervention, but not too early to assess the development of the relationship between therapist and patient. We use a patient rating of alliance based on the CALPAS scale.

- Total CALPAS scores (ranging from 0, indicating low alliance, to 7, indicating high alliance) were used in some of the analyses reported previously, but here we also use a binary alliance variable (1 if CALPAS score $\geq 5$, otherwise 0).

- Not measured in the control group.

Lewis et al, BJP (2002); Tarrier et al BJP (2004); Dunn & Bentall, Stats in Medicine (2007); Emsley, Dunn and White, Stats Methods in Medial Research (2010).
Example: missing data in SoCRATES

- 182 (88.3%) out of 207 patients in the treated groups provided data on the number of sessions attended. 56 patients from the CBT group and 58 from the SC group completed CALPAS forms at session 4 (overall 55.34%).

- The analysis here is based on all control participants but only those from treated groups who provide both a CALPAS and a record of the number of sessions.

- There were N=13 participants who didn’t attend sufficient sessions to have their therapeutic alliance assessed – potential bias here.

Lewis et al, BJP (2002); Tarrier et al BJP (2004); Dunn & Bentall, Stats in Medicine (2007); Emsley, Dunn and White, Stats Methods in Medial Research (2010).
Principal stratification in SoCRATES

- We can postulate the existence of two principal strata:
  
  - **High alliance** participants – those observed to have a high alliance in the therapy group together with those in the control group who would have had a high alliance had they been allocated to receive therapy.
  
  - **Low alliance** participants – those observed to have a low alliance in the therapy group together with those in the control group who would have had a low alliance had they been allocated to receive therapy.

Lewis et al, BJP (2002); Tarrier et al BJP (2004); Dunn & Bentall, Stats in Medicine (2007); Emsley, Dunn and White, Stats Methods in Medial Research (2010).
Mplus input: SoCRATES alliance

TITLE: Principal stratification – SoCRATES

DATA: FILE IS Socrates_alliance.raw;

VARIABLE: NAMES logdup pantot pant18 yearsed c1 c2 rgroup alliance resp;
CLASSES C(2);
CATEGORICAL are alliance resp;
USEVARIABLES logdup pantot pant18 yearsed c1 c2 rgroup alliance resp;
MISSING are pant18(999) alliance(999);

ANALYSIS: TYPE=MIXTURE;
STARTS = 100 10;
Mplus input: SoCRATES alliance

MODEL: %OVERALL%
resp ON logdup pantot yearsed c1 c2 rgroup;  !Missing data model
pant18 ON logdup pantot yearsed c1 c2 rgroup;  !Outcome model
C#1 ON logdup pantot yearsed c1 c2;  !Class model

%C#1% ! Low Alliance
[alliance$1@15];  !threshold to force alliance=0 into this class
[resp$1];  !release equality constraints on relevant model
resp ON rgroup*0;  !intercept terms for the effects of randomised
[pant18];  !intervention
pant18 ON rgroup*0;

%C#2% ! High alliance
[alliance$1@-15];  !threshold to force alliance=1 into this class
[resp$1];
resp ON rgroup*0;
[pant18];
pant18 ON rgroup*0;
Example: SoCRATES - results

**Estimated ITT effect on 18 month PANSS scores**

<table>
<thead>
<tr>
<th></th>
<th>Low alliance</th>
<th>High alliance</th>
</tr>
</thead>
<tbody>
<tr>
<td>Missing data ignorable (MAR)</td>
<td>+7.50 (8.18)</td>
<td>-15.46 (4.60)</td>
</tr>
<tr>
<td>Missing data ignorable (MAR)</td>
<td>0 (*)</td>
<td>-12.73 (4.75)</td>
</tr>
<tr>
<td>Missing data latently ignorable (LI)</td>
<td>+6.49 (7.26)</td>
<td>-16.97 (5.95)</td>
</tr>
<tr>
<td>Missing data latently ignorable (LI)</td>
<td>0 (**)</td>
<td>-13.50 (5.31)</td>
</tr>
</tbody>
</table>

* Zero ITT constraint in low alliance group (exclusion restriction)
** Compound exclusion restriction i.e. no ITT effect on PANSS or probability of missing value

Emsley, Dunn and White, Stats Methods in Medial Research (2010).
Example: PRP Trial aims & structure

- **Psychological Prevention of Relapse in Psychosis**
  Philippa A. Garety, David G. Fowler, Daniel Freeman, Paul Bebbington, Graham Dunn and Elizabeth Kuipers

  - Evaluation of Cognitive Behaviour Therapy (CBT) and Family Intervention (FI) for relapse prevention and reduction of positive symptoms in psychosis.
  
  - Aimed to test how CBT and FI work, based upon the cognitive model of psychosis (specific hypotheses concerning different mediators for CBT and FI).
  
  - Two parallel trials (pathways): one for people with carers (CBT vs. FI vs. TAU) and the other for those without (CBT vs. TAU).
  
  - Treatment trial accompanied by a series of theoretical studies of delusions and hallucinations.
PRP Trial: Summary of findings

• Primary outcomes:
  ➢ no ITT effects on recovery, relapse or readmission

• Secondary outcomes:
  ➢ only one significant effect of CBT (reduced depression (BDI) at 24 months). No effect on PANSS scores, for example, or putative mediators.
  ➢ no significant effects of FI.

This is not very promising!

But from further exploratory analyses there was a suggestion that CBT worked for participants with carers (moderator effect).

PRP Trial: Was treatment as intended?

- We consider patient engagement in therapy procedures as a potential treatment-effect moderator (for this we now drop the FI arm).

- Determined by careful examination of recordings of therapy sessions using the **Cognitive Therapy for Psychosis Adherence Scale (CTPAS)** and **Cognitive Therapy Scales (CTS)**.

- CTPAS/CTS classification:
  - No dose (21)
  - Medium dose (39)
  - Full dose (42)

- Dose not defined (missing) in TAU arm.

- A few dose assessments missing in CBT arm.

### PRP Trial: PANSS outcome at 12m & 24m

<table>
<thead>
<tr>
<th></th>
<th>Panss 0m</th>
<th>Panss 12m</th>
<th>Panss 24m</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>No dose</strong></td>
<td>63.0</td>
<td>56.4</td>
<td>52.3</td>
</tr>
<tr>
<td><strong>Medium dose</strong></td>
<td>66.2</td>
<td>60.1</td>
<td>58.9</td>
</tr>
<tr>
<td><strong>Full dose</strong></td>
<td>63.5</td>
<td>56.0</td>
<td>56.3</td>
</tr>
<tr>
<td><strong>Controls</strong></td>
<td>65.0</td>
<td>58.5</td>
<td>58.5</td>
</tr>
</tbody>
</table>

A high PANSS score implies a worse symptom outcome.

**This is still not very promising – high dose worse than no dose!**

* Note that only about 50% of the No dose group provide outcome data.

PRP Trial: Principal stratification

Defined as before in terms of potential response to randomisation

- **Statum 1**: a group of participants who receive little or no therapy whatever their treatment allocation.

- **Stratum 2**: a group of participants who would receive no therapy if allocated to the control condition but a medium dose of CBT if allocated to the treatment group.

- **Stratum 3**: a group who would receive no therapy if allocated to the control condition but a full dose of CBT if allocated to the treatment group.

- Membership of one of these three classes (the Principal Strata) is directly observable in the CBT arm but remains latent (hidden) under TAU.

- Principal stratum membership is independent of treatment allocation.

- Potentially, we can stratify by stratum membership and evaluate the ITT effects of treatment allocation within these strata.

Mplus input: PRP example

VARIABLE:
  NAMES ARE r1 r2 r3 r4 sex outpat
  pan0 pan12 pan24 resp12 treat c1 c2 c3;
  CATEGORICAL resp12;  ! Non-missing value indicator
  CLASSES c(3);  ! Engagement status
  TRAINING c1 c2 c3;  ! Uses observed engagement
  MISSING pan12 (999) pan24 (999);  ! as training data

  ....etc.

ANALYSIS:
  TYPE=MIXTURE
  ESTIMATOR=ML;
  STARTS = 1000 20;  ! Local maxima are a potential
  BOOTSTRAP=250;  ! problem when > 2 classes
MODEL: %OVERALL%
  pan12 ON treat carer r1 r2 r3 r4 pan0 sex outpat;
  C#1 ON carer r1 r2 r3 r4 pan0 sex outpat;
  C#2 ON carer r1 r2 r3 r4 pan0 sex outpat;

%C#1% ! None
[pan12];
pan12;
pan12 ON treat@0 ;

%C#2% ! Basic therapy
[pan12];
pan12;
pan12 ON treat*0;

%C#3% ! Full therapy
[pan12];
pan12;
pan12 ON treat*0;
PRP Trial: ITT estimates of effects of treatment as intended (6 runs in Mplus v6.1)

<table>
<thead>
<tr>
<th></th>
<th>No</th>
<th>Medium</th>
<th>Full</th>
</tr>
</thead>
<tbody>
<tr>
<td>Panss 12 months</td>
<td>0*</td>
<td>+6.4 (3.7)</td>
<td>-16.4 (6.8)</td>
</tr>
<tr>
<td>Panss 24 months</td>
<td>0*</td>
<td>+7.5 (4.6)</td>
<td>-11.3 (5.9)</td>
</tr>
<tr>
<td>BDI 12 months</td>
<td>0*</td>
<td>+2.7 (4.6)</td>
<td>-2.7 (4.6)</td>
</tr>
<tr>
<td>BDI 24 months</td>
<td>0*</td>
<td>+3.3 (3.4)</td>
<td>-7.6 (4.4)</td>
</tr>
<tr>
<td>Months remission 0-12</td>
<td>0*</td>
<td>-1.7 (1.1)</td>
<td>+5.6 (2.7)</td>
</tr>
<tr>
<td>Months remission 12-24</td>
<td>0*</td>
<td>-2.1 (1.4)</td>
<td>+2.1 (2.1)</td>
</tr>
</tbody>
</table>

*Exclusion restriction (constraint) with bootstrap standard errors

PRP Trial: marginal ITT estimates over 12m and 24m (Mplus v6.1)

Joint analysis of PANSS 12m & PANSSm 24 to get common ITT estimates.

<table>
<thead>
<tr>
<th></th>
<th>No</th>
<th>Medium</th>
<th>Full</th>
</tr>
</thead>
<tbody>
<tr>
<td>Estimate</td>
<td>0*</td>
<td>+5.2 (3.2)</td>
<td>-12.7 (4.1)</td>
</tr>
<tr>
<td>Estimate</td>
<td>0*</td>
<td>0*</td>
<td>-12.2 (4.8)</td>
</tr>
</tbody>
</table>

* Exclusion restriction (constraint)
With bootstrap standard errors

This assumes a Missing at Random missing data mechanism.

Extending analysis to repeated measures on the outcome variable

Randomisation → Mediator → Outcome → Outcome → Outcome

Time 1  Time 2  Time 3

Growth curves for repeated outcome measures

- This is the random part of a random coefficient model, which is the same as a traditional linear growth curve model.

- Occasion i, subject j

\[ y_{ij} = \beta_1 + \beta_2 x_{ij} + \zeta_{1j} + \zeta_{2j} x_{ij} + \epsilon_{ij} \]

\( y_{ij} \) is the outcome for subject j at occasion i. The model consists of:

- **Fixed part**: \( \beta_1 + \beta_2 x_{ij} \)
- **Random part**: \( \zeta_{1j} + \zeta_{2j} x_{ij} + \epsilon_{ij} \)
Example: SoCRATES trial again

- Instead of simply analysing the 18 month outcomes, we use the fact that the PANSS was administered at various time points:
  - at baseline (time score 0)
  - 6 weeks (1.94591)
  - 3 months (2.5649493)
  - 9 months (3.6109178)
  - 18 months (4.3694477)

- In the analyses we log transformed the timescale measured in weeks, and exploring each trajectory suggests a quadratic trajectory slope. We use Mplus v6.12.
Observed trajectories for 30 patients
Model fitted quadratic curves for 30 same patients
Principal strata with growth curves

High Alliance Class

Outcome 1, Outcome 2, Outcome 3

I → S

Low Alliance Class

Outcome 1, Outcome 2, Outcome 3

I → S
Mixture modelling

- Mixture modeling refers to modeling with categorical latent variables that represent subpopulations where population membership is not known but is inferred from the data – such as principal strata.

- The simplest longitudinal mixture model is latent class growth analysis (LCGA). In LCGA, the mixture corresponds to different latent trajectory classes. No variation across individuals is allowed within classes.

- Another longitudinal mixture model is the growth mixture model (GMM). In GMM, within class variation of individuals is allowed for the latent trajectory classes. The within-class variation is represented by random effects, that is, continuous latent variables, as in regular growth modeling.
Growth mixture model

The \( \varepsilon \) represent measurement error and time specific variation.

\( \varepsilon_1, \varepsilon_2, \varepsilon_3, \varepsilon_4, \varepsilon_5 \)

\( Y_1, Y_2, Y_3, Y_4, Y_5 \)

I – zero time score for the slope growth factor at \( t=1 \) defines the intercept growth factor as an initial status factor.

All random effect means are specified as varying across latent classes.

Latent class membership is predicted as a function of baseline covariates.

The treatment effect is captured by a regression of the linear and quadratic slopes on random allocation, and is allowed to vary across latent classes.
SoCRATES analysis in Mplus v6.12

- We simultaneously fit the following models using ML with the EM algorithm:
  - Principal strata membership on covariates (log of duration of untreated psychosis, centre, years of education).
  - Quadratic growth curve model within each class, allowing all the random effect means and variances to vary between high and low alliance classes.
  - Effect of randomisation on the slope within each class.

- Bootstrap the procedure to obtain valid standard error estimates.

- Missing data under MAR allowed for outcomes.
Mplus input: SoCRATES growth models

Data:
  File is SoCRATES_Growth_models.dat ;
Variable:
  Names are id therapy logdup pantot pant18 sfsbase centre cptot4 sessions yearsed cbt sc c1 c2 group cpmax lgp c1gp c2gp yrgp pgp sessbin alliance s_a pan1 pan3 pan9;
  Missing are all (999) ;
  CLASSES C(2);  
  CATEGORICAL alliance;
  USEVAR are pantot pan1 pan3 pan9 pant18 logdup yearsed c1 c2 alliance group;

Analysis:
  Type = MIXTURE;
  STARTS = 100 10; 
  ESTIMATOR=ml; 
  BOOTSTRAP=250;
MODEL: %OVERALL%

C#1 ON logdup yearsed c1 c2;
I S Q | pantot@0 pan1@1.94591 pan3@2.5649493 pan9@3.6109178 pant18@4.3694477;

S ON group;

%C#1% ! Low Alliance
[alliance$1@15];
I; S;
S ON group;

%C#2% ! High alliance
[alliance$1@-15];
I; S;
S ON group;
SoCRATES: Estimated means for low alliance class (N=63) and observed trajectories
SoCRATES analysis in Mplus v6.12

Latent Class 1 – Low Alliance Group (N=63)

Effect of Randomisation on SLOPE

Coeff=+1.808      SE=1.644      T=1.100      P-value=0.271

Random Effect Means/Intercepts

<table>
<thead>
<tr>
<th></th>
<th>Coeff</th>
<th>SE</th>
<th>T</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>INTER</td>
<td>90.444</td>
<td>3.441</td>
<td>26.281</td>
<td>0.000</td>
</tr>
<tr>
<td>SLOPE</td>
<td>-17.118</td>
<td>2.697</td>
<td>-6.346</td>
<td>0.000</td>
</tr>
<tr>
<td>QUADRATIC</td>
<td>2.269</td>
<td>0.460</td>
<td>4.928</td>
<td>0.000</td>
</tr>
</tbody>
</table>
SoCRATES: estimated means for high alliance class (N=138) and observed trajectories
SoCRATES analysis in Mplus v6.12

Latent Class 2 – High Alliance Group (N=138)

Effect of Randomisation on SLOPE

Coeff=-2.843      SE=1.136      T=-2.502      P-value=0.012

Random Effects Means/Intercepts

<table>
<thead>
<tr>
<th>Mean</th>
<th>Coeff</th>
<th>SE</th>
<th>T</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>INTER</td>
<td>87.844</td>
<td>1.954</td>
<td>44.955</td>
<td>0.000</td>
</tr>
<tr>
<td>SLOPE</td>
<td>-11.558</td>
<td>1.763</td>
<td>-6.556</td>
<td>0.000</td>
</tr>
<tr>
<td>QUADRATIC</td>
<td>1.784</td>
<td>0.309</td>
<td>5.769</td>
<td>0.000</td>
</tr>
</tbody>
</table>
Sample and estimated means by class
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Estimating causal parameters using parametric regression models

- Previous work on identification and estimation of direct and indirect causal effects using parametric regression models in VanderWeele and Vansteelandt (2009, 2010).
  - Outcomes can be continuous, binary, count.
  - Mediators can be binary or continuous.

- The traditional Baron and Kenny approach doesn’t allow for the presence of exposure-mediator interactions in the mediation analysis model.

- Causal mediation methods now extended to allow for interactions as well.

- Implemented in SAS, SPSS, R, Stata...
New Stata command: \texttt{PARAMED}

- \texttt{paramed} allows continuous, binary or count outcomes, and continuous or binary mediators, and requires the user to specify an appropriate form for the regression models.

- \texttt{paramed} provides estimates of the controlled direct effect, the natural direct effect, the natural indirect effect and the total effect with standard errors and confidence intervals derived using the delta method by default, with a bootstrap option also available.

Causal mediation analysis in Mplus

- Muthén (2011) presented causally-defined direct and indirect effects for:
  - Continuous, binary, ordinal, nominal and count variables.
  - New extension to mediation by a nominal variable.
  - Sensitivity analysis.

- MODEL CONSTRAINT is used to specify the causal direct and indirect effects, computed by specifying NEW parameters.

- Simple example: continuous mediator, continuous outcome, treatment-mediator interaction.
If the model is:

\[ \begin{align*}
    m_i &= \alpha_0 + \alpha_1 x_i + \alpha_2 l_i + \epsilon_{1i} \\
    y_i &= \beta_0 + \beta_1 x_i + \beta_2 m_i + \beta_3 x_i m_i + \beta_4 c_i + \beta_5 l_i + \epsilon_{2i}
\end{align*} \]

Then, applying the formal definitions, under the appropriate assumptions,

\[ \begin{align*}
    CDE(m) &= \beta_1 + \beta_3 m \\
    PNDE &= \beta_1 + \beta_3 \alpha_0 \\
    TNIE &= \beta_2 \alpha_1 + \beta_3 \alpha_1
\end{align*} \]
Mplus input: Monte Carlo simulation of y m x xm

model:
[y*1] (beta0); !intercept
y on x*.4 (beta2);
y on xm*.2 (beta3);
y on m*.5 (beta1);
[m*2] (gamma0); !intercept
m on x*.5 (gamma1);
y*.5; !residual variance
m*1; !residual variance

model constraint:
new(tie*.35 pie*.25 de*.8);
tie=beta1*gamma1+beta3*gamma1;
pie=beta1*gamma1;
de=beta2+beta3*gamma0;
## Muthén (2011) Table 1

### MODEL RESULTS

<table>
<thead>
<tr>
<th>ESTIMATES</th>
<th>S. E.</th>
<th>M. S. E.</th>
<th>95%</th>
<th>% Sig</th>
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<td><strong>Population</strong></td>
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<td><strong>Std. Dev.</strong></td>
<td><strong>Average</strong></td>
<td><strong>Cover Coeff</strong></td>
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<td>ON</td>
<td>X</td>
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<td><strong>Intercepts</strong></td>
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<tr>
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<td>0.0702</td>
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<td><strong>New/Additional Parameters</strong></td>
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<td>TIE</td>
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<td>DE</td>
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<td>0.8027</td>
<td>0.0802</td>
<td>0.0766</td>
</tr>
</tbody>
</table>
Contents

1. Discussion of mediation and statistical mediation analysis

2. Causal mediation parameters

3. Process evaluation (using Mplus)

4. Causal mediation analysis (using Mplus)

5. Conclusions
Some conclusions

- Statistical mediation (B&K) has three main problems:
  1. Unmeasured confounding
  2. No interactions between exposure and mediator on outcome
  3. Doesn’t include non-linear models

- Causal mediation analysis has arisen from the causal inference literature, and addressed these problems.

- Available in other software, and now also in Mplus thanks to Muthén (2011):
  - Can be applied to new settings (nominal mediators)
  - But overall slightly cumbersome?
Some conclusions (2)

• Principal stratification can be used to analyse process variables, with singly observed and repeated measures of outcomes. Can extend to multiple classes/strata.

• Real strength of Mplus is the longitudinal modelling features, and potential for combining this with mediation analysis in a mixture framework.

• Not currently explored elsewhere;
  ➢ Extending definition of PNDE and TNIE etc. to longitudinal data?
Selected references


Research Programme: Efficacy and Mechanisms Evaluation

Funded by MRC Methodology Research Programme

- **Design and methods of explanatory (causal) analysis for randomised trials of complex interventions in mental health (2006-2009)**
  - Graham Dunn (PI), Richard Emsley, Linda Davies, Jonathan Green, Andrew Pickles, Chris Roberts, Ian White & Frank Windmeijer.

- **Estimation of causal effects of complex interventions in longitudinal studies with intermediate variables (2009-2012)**
  - Richard Emsley (MRC Fellow), Graham Dunn.

- **Designs and analysis for the evaluation and validation of social and psychological markers in randomised trials of complex interventions in mental health (2010-12)**
  - Graham Dunn (PI), Richard Emsley, Linda Davies, Jonathan Green, Andrew Pickles, Chris Roberts, Ian White & Frank Windmeijer with Hanhua Liu.

- **PhD students:** Lucy Goldsmith and Clare Flach (2010 – 2013)
  Philip Foden (2011-2014)